

Characterizing individual differences in animal responses to a nutritional challenge: Toward improved robustness measures

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ABSTRACT

The aim of this paper was to explore the variation between individuals in the response to and recovery from a nutritional challenge, the repeatability of responses between lactation stages, and the use of shape-clustering methods to classify similar individuals. Sixteen dairy goats were exposed to a 2-d nutritional challenge (underfeeding) at 2 different stages of lactation. Each challenge consisted of a 7-d control period with standard total mixed ration (TMR), 2 d of straw-only feeding, and a 10-d recovery period on the TMR. All feeds were offered ad libitum, as was water. The first challenge was in late lactation on primiparous goats (mean days in milk = 249), and the second challenge was carried out on the same goats early in the following lactation (mean days in milk = 28). The main energetic response traits dry matter intake (DMI), milk yield, body weight, milk fat and protein contents, and plasma glucose, fatty acids (NEFA), β-hydroxybutyrate (BHB), urea, and insulin, were measured daily throughout. A clustering procedure linked to a piecewise mixed model was used to characterize different types of response. As expected, straw feeding caused a large decline in DMI and milk yield, and substantial increases in milk fat and milk protein composition, relative to the prechallenge period on the control TMR. For both DMI and milk yield, the slope of the response, and hence the size of the drop, was strongly related to the prechallenge values, indicating that these 2 measures were tightly constrained by the challenge. Regression slopes between lactation stages for responses to the same nutritional challenge were significant for DMI, milk protein content, plasma BHB and urea, and body weight, indicating that within-animal responses in late and early lactation were repeatable. The clustering procedure generally performed well, classifying both scaling differences and differences in shape. The extent of reranking of cluster designations between late lactation and the following early lactation period was examined. For milk yield, DMI, body weight, and urea, relatively little reranking occurred; the numbers of goats not changing class number were 10, 12, 10, and 13, respectively. In contrast, for milk contents of fat and protein, and also for BHB, no clear association was found between late and early lactation class numbers. For NEFA and glucose, these comparisons were not relevant because either the vast majority of goats were in 1 cluster (NEFA) or because an outlier goat skewed the cluster designation (glucose in late lactation). For insulin, 9 out of 16 goats kept the same rank.

Key words: dairy goat, clustering method, adaptive capacity, resilience, lactation

INTRODUCTION

The ability of an animal to respond to, and recover from, environmental challenges is an increasingly important trait. This adaptive capacity is a key component of animal robustness [the ability to maintain life functions in the face of constraining environments (Kitano, 2004)] in the context of the challenges facing future livestock production. The diversity of environments to which livestock will be exposed is expected to increase as increasing food demand leads to further exploitation of marginal land (Bocquier and Gonzáles-Garcia, 2010). In addition, climate change is driving an increase in the variability over time in environmental conditions with a higher frequency of extreme conditions (Hansen et al., 2012). These 2 trends will place increasing demands on the animal's adaptive capacity.

Although there is now general agreement that adaptive capacity is a complex trait conferred by a combination of several underlying components (Strandberg, 2009), there is little agreement on how to quantify it, and on which are the key biological components (Friggens et al., 2010). This clearly presents a difficulty for characterizing both adaptive capacity and thus robust-

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ness. Marked differences are present between breeds and strains in performance responses to environmental challenges (Horan et al., 2005; Delaby et al., 2009). It has also been shown that individual variation in performance, and in the use of body reserves, can be adversely affected in harsh environments (Puillet et al., 2010). In this context, being able to characterize robustness at the level of the individual animal would be valuable for refining both management and selection strategies with respect to the anticipated increase in diversity of livestock systems.

The present study fits within an overall aim to develop ways to characterize the robustness and adaptive capacity of animals that can then be adapted to become operational methods for quantifying robustness on farm. [Robust animals are those that are able to be healthy and perform well under a wide range of environmental conditions (Amer, 2012)]. We believe that this is a timely and realistic aim given the increasing panel of automated measures that are becoming available on farm in the context of precision livestock farming (Rutten et al., 2013). Such measures provide the opportunity to describe the dynamic features of responses of individual animals to environmental perturbations (Codrea et al., 2011). The advances in on-farm technology also mean that such response dynamics can be measured across multiple performance and physiological traits. Preliminary studies indicate that it is possible to derive operational definitions of components of robustness that are quantifiable. For example, it has been shown that biological entities such as degree of infection can be derived from multivariate measures of the animal's disease response (Højsgaard and Friggens, 2010).

In the area of nutritional status and nutritional challenges, markers to monitor key metabolites such as BHB have been in use for a considerable time and now exist in automated systems (Nielsen et al., 2005). Several indexes are also now available for the purpose of monitoring to identify at-risk animals with respect to health status (Bramley et al., 2008; Moyes et al., 2013). In the present study, we wanted to extend this approach from a monitoring perspective toward phenotyping (i.e., to develop ways to better understand and characterize between animal variation in adaptive capacity using multivariate measures).

Accordingly, an experiment was carried out in which time-series measurements of behavior, performance, and metabolites were made in dairy goats exposed to a nutritional challenge, at 2 different stages of lactation. A major issue of this work relates to the exploration of the resulting data because the focus of this work is not to quantify the overall effect of the nutritional challenge but rather to examine individual differences

in responses. Thus, we are interested in the portion of the variation that is usually eliminated (i.e., assigned to the error term) in standard ANOVA approaches. Given this, and the fact that there was no strong a priori rationale for which measures (or aspects of measures) are key components of adaptive capacity, an exploratory data analysis was used in this paper. Therefore, the aim of this paper was to prepare the ground (proof of concept) by providing the methodological basis for quantifying biologically meaningful descriptions of adaptive capacity phenotypes from multivariate measures. More specifically, the aim was to develop a methodology to explore the variation between individuals in the response and recovery to the challenge, the repeatability of responses, and shape-clustering methods to classify similar individuals.

MATERIALS AND METHODS

Animals, Feeds, and Design

Sixteen dairy goats, housed in individual pens (1.2 m by 0.75 m), were exposed to a 2-d nutritional challenge (underfeeding), at 2 different stages of lactation. Each challenge consisted of a 7-d control period with standard TMR, 2 d of straw only feeding, and a 10-d recovery period on the TMR. Prior to the start of each challenge period, the goats had received the standard TMR for at least 15 d. All feeds were offered ad libitum, as was water. The first challenge was in late lactation (mean DIM = 249, SD = 2.8, all goats were primiparous), and the second challenge was carried out on the same goats early in the following lactation (mean DIM = 28, SD = 3.3). The goats were housed in individual pens each with their own feed trough; they were milked twice per day and feed was distributed twice daily shortly after milking. Animals were cared for and handled in accordance with the French legislation on animal experimentation and European Convention for the Protection of Vertebrates Used for Experimental and Other Scientific Purposes (European Directive 86/609). The experiment was carried out between October 12, 2009, and February 14, 2010.

The standard TMR (on a DM basis) consisted of 20% chopped hay, 30% chopped dried alfalfa (Rumiluz, Désialis, Paris, France), 30% sugar beet pulp, and 20% of a commercial dairy concentrate (containing 18% maize, 14% sugar beet pulp, 12% sunflower meal, 10% wheat, 10% soybeans, 9% rapeseed meal, 6% soybean meal, 4% wheat distillers grains, 3.5% linseed, 3% pea seed, 1% rapeseed oil, 3% molasses, and 6.5% mineral and vitamin premix). The standard TMR (DM content 98.3%) had a measured content (on a DM basis) of 12.0% CP, 5.6% starch, 8.9% ash, 37.2% NDF, 20.0%

ADF, and 2.5% ADL. The chopped straw (DM content 97.2%) contained (on a DM basis) 2.5% CP, 8.4% ash, 74.1% NDF, 43.1% ADF, and 5.1% ADL.

Measurements

From the beginning of the prechallenge control period until the end of the postchallenge recovery period measurements of feed intake and BW were made daily. Milk yield was recorded and milk samples were taken at each milking for standard milk composition measures (Fossomatic, Hillerød, Denmark). Blood samples were taken daily in heparinized tubes after the morning milking and before feeding (average 20 min pre-feeding), immediately centrifuged for 10 min at $3,000 \times q$ at 4°C, and the plasma was frozen (-20°C) for subsequent analysis, in duplicate, of urea (intra- and interassay CV: 1.7 and 3.2%), glucose (intra- and interassay CV: 0.8 and 1.8%), BHB (intra- and interassay CV: 1.4 and 2.0%), and fatty acids (NEFA; intra- and interassay CV: 4.1 and 8.2%) using a Cobas Mira-Analyzer (Roche, Mannheim, Germany) with commercial kits for urea (urea/BUN, Roche), glucose (Gluco-quant, Glucose/ HK, Roche), NEFA [NEFA-HR(2), Wako Chemical GmBH, Neuss, Germany, and BHB by the procedure of Barnouin et al. (1986). Insulin was measured using an ELISA kit (10-1202-01; Mercodia AB, Uppsala, Sweden; intra- and interassay CV: 1.2 and 4.5%).

Feeds were analyzed using the following standard methods: DM estimated from water content (ISO, 1983), ash (ISO, 1978), and starch (ISO, 2004). Total N was determined by the Dumas technique (Sweeney and Rexroad, 1987). Cell wall content was estimated by the NDF method of Van Soest and Wine (1967), modified by Giger et al. (1987) with the use of a heat-stable α -amylase but without sodium sulfite and decalin. Lignocellulose (ADF) and lignin (ADL) were obtained using a sequential approach on the NDF residue (Giger et al., 1987). All the cell wall components were expressed on an ash-free basis.

Statistical Analyses

In the present study, we focused on the main energetic response traits (i.e., DMI, milk yield, BW, milk fat and protein contents, and plasma glucose, NEFA, BHB, urea, and insulin). The statistical analysis was carried out on each trait separately and consisted of 2 steps. In the first step, the response profile of each goat was described using a piecewise model with 4 parameters. These parameters that characterize the response recovery profile allowed us to describe (1) the relationship between prechallenge and response to challenge, and

(2) the relationship between response to challenge and recovery. In a second step, all response profiles were jointly analyzed to identify cluster of animals exhibiting similar profiles.

Step 1: Individual Analyses. For each trait, and each lactation stage (late vs. early), each of the 16 individual time series were characterized separately using the following piecewise model:

$$\begin{split} Y_t &= V_1 \times I_{(t \leq 0)} + V_2 \times t \times I_{(0 < t \leq 2)} + \\ & (V_3 \times t + V_4 \times t \times 2) I_{(2 < t < 4)} + V_5 \times I_{(t > 4)} + E_t, \end{split}$$

where Y_t is the trait measured at time t and the coefficients V_n are estimated by the model, V_1 is the prechallenge level (assumed to be constant throughout the prechallenge period), V_2 is the linear slope of the response during the 2-d challenge, V_3 is the linear component of the recovery, V_4 is the quadratic component (deceleration) of the recovery, and V_5 is the postchallenge level that was assumed to be complete and constant 6 d after the beginning of challenge (confirmed by visual inspection). The error terms E_t are assumed to be Gaussian and independent. Indicator variable $I_{\{condition\}}$ is a dummy variable with value 1 if condition is true, and 0 otherwise. The final postchallenge level was not assumed to be the same as the prechallenge level. In addition, parameters V_1 to V_5 were estimated subject to the constraint of profile continuity. The model and its associated parameters are summarized in Figure 1. The choice of model, along with its inherent assumptions, was made on the basis of both a visual inspection of the response/recovery profiles and keeping in mind a parsimony principle (keeping an acceptable ratio between the number of parameters and the number of data points in an individual time series). The set of parameters V_1 to V_5 was then used to characterize the profile of each goat.

The relationship between prechallenge and response to challenge was analyzed using the following model:

$$\begin{split} V_{2ij} &= \mu + a \times LA\,CT_i + b \times \tilde{V}_{1ij} + c \times LA\,CT_i \times \tilde{V}_{1ij} \\ &+ G_j + E_{ij}, \end{split}$$

where V_{2ij} is coefficient V_2 for goat j and lactation stage i, \tilde{V}_{1ij} is the mean centered coefficient V_1 [i.e., $\tilde{V}_{1ij} = V_{Iij} - mean(V_{Ii})$]. Parameter μ is the intercept for early lactation, $LACT_i$ is a dummy variable with values 0 and 1 for early and late, respectively, a is the coefficient for difference in intercepts between lactation stages, b is the slope coefficient measuring the link between V_1 and V_2 , and v_3 is the coefficient for difference in slopes be-

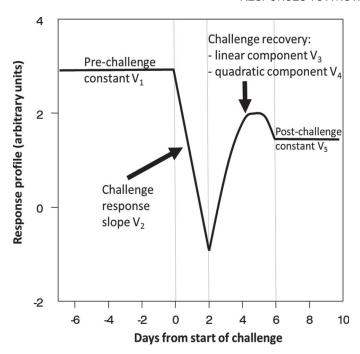


Figure 1. Schematic representation of the piecewise model for describing response and recovery time-series profiles relative to a nutritional challenge.

tween lactation stages. Last, G_j is a Gaussian random effect with variance σ_G^2 accounting for possible correlations between coefficients measured on the same goat, and E_{ii} is the residual with variance σ^2 . All random effects are assumed independent. This model was applied to each of the following traits: DMI, milk yield, milk fat content, milk protein content, BW, plasma glucose, plasma NEFA, plasma BHB, plasma urea, and plasma insulin. As across all traits c was found not to be significant, this interaction term was dropped from the model.

The relationship between response to challenge and recovery was analyzed using a model similar to model (1), with the following differences:

- The variable to be explained is now the rate of
- deceleration in recovery V_{4ij} ,

 The explanatory variable is now $\tilde{V}_{2ij} = V_{2ij}$ $mean(V_{2i}).$

Note that V_4 is the quadratic component of the recovery and characterizes (half) the rate of deceleration in the recovery. This variable was chosen because the absolute rate of recovery (i.e., the combination of the quadratic and linear components of the recovery) is strongly correlated with the amplitude of the response (r > 0.8 for most traits). One animal in late lactation was excluded as an extreme outlier. As previously, c was found not to be significant across all traits; therefore, this interaction term was dropped from the model.

To assess repeatability between lactation stages, within trait, the following linear regression was used:

$$V_{22j} = a + b \times V_{21j} + E_j.$$

Parameters a and b are the intercept and slope for the relation, respectively. Note that perfect repeatability would result in b = 1. A similar analysis was performed

Step 2: Clustering Analysis. To identify goats exhibiting a similar response profile to the challenge, a clustering analysis was performed based on a 2-stage model. Each goat is assumed to belong to one of the Kclusters, where the number of clusters K is fixed and assumed to be known. We note π_k as the proportion associated with class k (i.e., the prior probability to belong to class k). For a goat j that belongs to cluster k, one has

$$\begin{split} Y_{jt} &= V_1^k \times I_{(t \leq 0)} + V_2^k \times t \times I_{(0 < t \leq 2)} \\ &+ (V_3^k \times t + V_4^k \times t \times 2) I_{(2 < t \leq 4)} + V_5^k \times I_{(t > 4)} + E_{jt}, \end{split}$$

which corresponds to model (1) previously described, except that parameters V_1^k, \ldots, V_5^k are common to all animals belonging to class k, and that the variance of E_{it} is assumed to be the same for all animals. The goal of the analysis is then to estimate the coefficients $(\pi_k, V_1^k, \ldots, V_5^k)$ but also the posterior probabilities τ_{ik} of each animal j to belong to class k. This can be done thanks to the expectation maximization (EM) algorithm (Dempster et al., 1977). In practice, the number of clusters K being unknown, it was chosen within the range of cluster numbers from 2 to 5 based on the integrated completed likelihood (ICL) criterion (Biernacki et al., 2000). Both the EM algorithm and the ICL criterion are described in the Appendix. To allow subsequent comparison of cluster class attributions between traits and lactation stages, the analyses were rerun using a pre-specified fixed number of 3 clusters, which corresponded to the cluster number most frequently found across traits.

Comparison across traits was carried out on the cluster classifications using the adjusted Rand index (ARI: Rand, 1971). This is a pair-counting-based method that calculates across any 2 traits whether the cluster classifications of a given pair of goats in the first trait match the cluster classifications of the same pair of goats in the second trait. When this is done across all possible pair combinations of goats, it provides an index of clustering agreement between traits that makes no a

priori assumptions about mapping of one cluster classification onto the other, and that is adjusted for the fact that purely random clustering would still create some matches. Using ARI, a perfect match would give an ARI = 1 and purely random matches would give an ARI = 0. In addition, principal components analysis (**PCA**) was carried out on the cluster classification membership. The procedure was the following: for each initial variable, the clustering score was obtained (with values 1, 2, and 3 corresponding to a clustering into 3 classes). This clustering score was then replaced by 3 dummy variables: Ind1, Ind2, Ind3, where $\operatorname{Ind}_{jk} = 1$ if

animal j belongs to cluster k, 0 otherwise; thus, no a priori relation was assumed between cluster classes. A PCA was performed on the set of all dummy variables corresponding to all traits, for each lactation stage.

RESULTS

The average time trends of DMI, milk yield, milk fat, and milk protein contents through the prechallenge, challenge, and postchallenge periods in both late and the following early lactation are shown in Figure 2. The

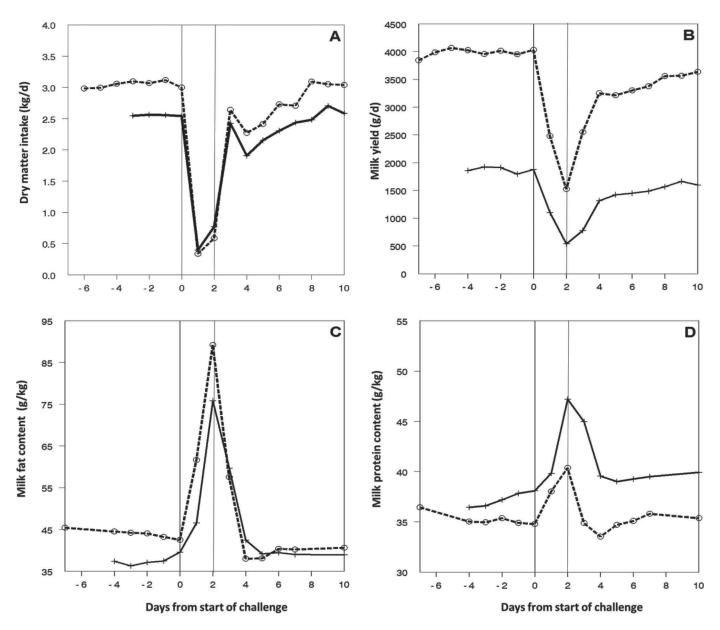


Figure 2. The average time trends of (a) DMI, (b) milk yield, (c) milk fat, and (d) milk protein contents through the prechallenge (before d 0), challenge (d 0 to 2), and postchallenge (d > 2) periods in both late (solid line) and the following early lactation (dashed line).

Table 1. Average prechallenge levels in late lactation and in the following early lactation¹

Item	Late lactation ²	Early lactation
DMI (kg/d)	2.48 (2.02-3.29; 0.36)	2.97 (2.40-3.60; 0.35)
Milk yield (kg/d)	1.93 (1.28–2.72; 0.46)	3.97(2.18-5.17; 0.65)
Milk fat content (g/kg)	35.7 (32.6–52.8; 2.9)	43.7 (36.5–49.4; 3.6)
Milk protein content (g/kg)	36.5 (30.9–45.8; 3.8)	35.3 (30.6–40.1; 2.4)
BW (kg)	60 (51–74; 7.4)	62 (54–78; 7.4)
Plasma glucose (g/L)	$0.60 \ (0.55-0.67; \ 0.03)$	0.57 (0.50-0.63; 0.04)
Plasma fatty acids (NEFA; μmol/L)	130 (97–149; 16)	493 (200–655; 136)
Plasma BHB (mg/dL)	32 (27–41; 3.9)	52 (33–77; 14.2)
Plasma urea (g/L)	0.25 (0.18 - 0.32; 0.04)	$0.\overline{23} \ (0.16 - 0.30; \ 0.04)$
Plasma insulin (µg/L)	0.48 (0.28–1.04; 0.19)	0.18 (0.08–0.44; 0.09)

¹The prechallenge levels were estimated within a piecewise mixed model (see Materials and Methods), the ranges followed by the standard deviations are given in parentheses.

average prechallenge values for all traits in late and early lactation are given in Table 1.

Response to a Nutritional Challenge (2 d Straw-Only Feeding)

The slopes of response (V_2 ; see Figure 1) to the change from a standard TMR to straw-only feeding in early and late lactation are presented in Table 2. As expected, straw feeding caused a large decline in DMI and milk yield, and substantial increases in milk fat and milk protein composition, relative to the prechallenge period on the control TMR. For both DMI and milk yield, the slope of the response, and hence the size of the drop, was strongly related to the prechallenge values indicating that these 2 traits were tightly constrained by the challenge. The responses in DMI expressed relative to pre-period values (V_2/V_1) were -0.76 and -0.88 in late and early lactation, respectively. The proportional

responses in milk yield were -0.75 and -0.62 in late and early lactation, respectively. In both cases, the differences between lactation stages were significant. For all other traits (milk fat, milk protein, NEFA, BHB, glucose, urea, insulin), the prechallenge values had no significant effect.

For intake, milk protein content, plasma BHB and urea, and BW, significant slope coefficients were observed from regressions relating the rates of response (V_2) between lactation stages, albeit with modest \mathbb{R}^2 (Table 3). This indicates that within-animal responses in late and early lactation were repeatable (Figure 3). For plasma metabolites other than BHB and urea, the relationship between lactation stages in rates of response (V_2) was not significant, indicating that the other metabolic adaptations to nutritional challenge were not repeatable within animal. With the exception of glucose, insulin, and milk protein content, all intercepts of these regressions were significant, indicating

Table 2. The relationship between the slope of the response during nutritional challenge, stage of lactation (early vs. late), and the prechallenge level of the measure¹

Response slopes $(V_2, \text{ units per day})$	Early lactation	Late lactation adjustment	Coefficient for prechallenge level (V_1)	Residual	R^2
DMI (kg/d)	-1.31	+0.38	-0.49	0.005	0.93
Milk yield (kg/d)	-1.25	+0.55	-0.44	0.007	0.95
Milk fat content (g/kg)	21.6	-3.5	+0.05	6.43	0.46
Milk protein content (g/kg)	2.34	+2.60	+0.11	3.54	0.49
BW (kg)	-3.41	+1.85	-0.02	0.19	0.87
Plasma glucose (g/L)	-0.027	-0.006	-0.150	0.002	0.00
Plasma fatty acids (NEFA; µmol/L)	845	-410	-0.3	26,600	0.61
Plasma BHB (mg/dL)	8.34	-8.95	+0.03	57.89	0.38
Plasma urea (g/L)	0.164	-0.041	-0.283	0.0005	0.66
Plasma insulin (µg/L)	-0.042	-0.081	-0.111	0.003	0.37

¹The response slopes (V_2) and prechallenge levels (V_1) were estimated using a piecewise mixed model (see Materials and Methods). Results are shown in equation form: V_2 = average early lactation value + late lactation adjustment + coefficient × V_1 . In the analysis, the prechallenge levels were centered (i.e., expressed as the difference from the mean prechallenge level) and thus the average values of the response slopes in early and late lactation are given directly in the table. Coefficients significantly different from 0 are shown in bold (P < 0.05). As no significant interactions were found between prechallenge level and stage of lactation for any of the measures, this term was excluded from the model.

²One extreme animal excluded.

The R^2 was calculated from the residual of the full model (res_t) and the residual from the null model (res_n) as: (res_n - res_t)/res_n.

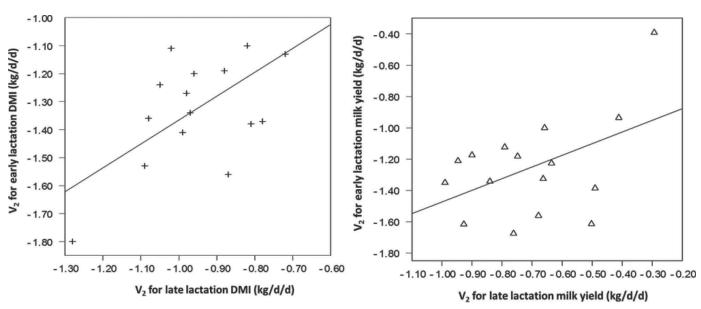


Figure 3. The relationship between late lactation and the following early lactation in slope of response to the same nutritional challenge (2 d of straw-only feeding) for (a) DMI and (b) milk yield.

that stage of lactation significantly affected the average size of response (Table 3).

Recovery from a Nutritional Challenge (2 d Straw-Only Feeding)

Given the inherently symmetrical nature of a response recovery profile, the size of the recovery (difference from peak to postchallenge baseline) is highly correlated with the size of response, consequently the stage of lactation effects on size of recovery (not shown) were essentially the same as those for size of response (Table 3). Thus, the quadratic component of the recovery $[V_4; i.e., (half)$ the rate of deceleration in the recovery] was focused on. As shown in Table 4, this rate (V_4) was

significantly greater in early lactation (relative to late lactation) for DMI, milk yield, and BW. Conversely, the rate of deceleration in recovery (V_4) was significantly less in early lactation for milk fat content, NEFA, BHB, urea, and insulin. The correlations between the rate of deceleration and the size of the response were above 0.5 for the majority of traits. This reflects the fact that within a similar recovery time, the greater the response, the "sharper" the curve needed to return to a plateau level. Indeed, the slope of response (V_2) had a significant effect on the rate of deceleration in recovery (V_4) for milk yield, BW, glucose, NEFA, BHB, and urea (Table 4). Thus, the relative rate of recovery/slope of response (V_4/V_2). This relative rate of recovery

Table 3. The relationship between early and late lactation values of the response slope during nutritional challenge derived by linear regression: response slope in early lactation $= a + b \times$ (response slope in late lactation)¹

Item	Coefficient a	Coefficient b	Residual	R^2adj
DMI (kg/d)	-0.51*	0.85	0.16	0.40
Milk yield (kg/d)	-0.73	0.74*	0.29	0.17
Milk fat content (g/kg)	13.47	0.45*	2.99	0.19
Milk protein content (g/kg)	0.81	0.31	1.65	0.23
BW (kg)	-2.34	0.68	0.61	0.38
Plasma glucose (g/L)	-0.027*	0.017	0.044	0.00
Plasma fatty acids (NEFA; μmol/L)	1,075	-0.53	211	0.00
Plasma BHB (mg/dL)	9.92	2.58	9.78	0.37
Plasma urea (g/L)	0.089	0.608	0.032	0.42
Plasma insulin (µg/L)	-0.042	0.0011	0.051	0.00

¹Perfect repeatability would result in b = 1. The response slopes (V_2) were estimated using a piecewise mixed model (see Materials and Methods). Coefficients significantly different from 0 are shown in bold (P < 0.05). *0.05 < P value < 0.07.

Table 4. The relationship between the rate of deceleration in recovery following nutritional challenge, stage of lactation (early vs. late), and the size of the response during the challenge¹

Rate of deceleration in recovery $(V_4, \text{ units per day}^2)$:	Early lactation	Late lactation adjustment	Coefficient for effect of response slope (V_2)	Residual	R^2
DMI (kg/d)	-0.241	0.049	0.133	0.0016	0.41
Milk yield (kg/d)	-0.165	0.102	0.140	0.0020	0.67
Milk fat content (g/kg)	5.87	-2.95	0.08	1.32	0.62
Milk protein content (g/kg)	1.02	-0.30	0.07	0.457	0.02
BW (kg)	-0.804	0.280	0.166	0.033	0.58
Plasma glucose (g/L)	-0.026	0.007	0.259	0.00009	0.47
Plasma fatty acids (NEFA; μmol/L)	287	-147	0.27	289.4	0.96
Plasma BHB (mg/dL)	5.45	-3.25	0.140	1.33	0.76
Plasma urea (g/L)	0.0756	-0.0139	0.2193	0.00006	0.57
Plasma insulin (µg/L)	-0.029	-0.044	0.079	0.00062	0.58

¹The rate of deceleration in recovery (the quadratic term of the recovery; V_4) and the size of the response (V_2) were estimated using a piecewise mixed model (see Materials and Methods). Results are shown in equation form: V_4 = average early lactation value + late lactation adjustment + coefficient × V_2 . In the analysis, the V_2 values were centered (i.e., expressed as the difference from the mean value), and thus the average values of the rate of deceleration in recovery in early and late lactation are given directly in the table. Coefficients significantly different from 0 are shown in bold (P < 0.05). As no significant interactions were found between V_2 and stage of lactation for any of the measures, this term was excluded from the model. The R^2 was calculated from the residual of the full model (res_f) and the residual from the null model (res_n) as: (res_n - res_f)/res_n.

was significantly greater than zero for DMI, milk yield, milk fat content, BW, plasma glucose, NEFA, and urea (results not shown). This indicates that for these traits the recovery was quicker than would be expected for the size of the initial response. No traits were found for which the recovery was more gradual than expected according to the size of the response. For milk yield, milk fat content, and BW, a significant difference was found between early and late lactation relative recovery rates. However, when examined for repeatability between late and early lactation, no significant effect was found on early lactation values of relative recovery rate in late lactation.

Clustering of Response-Recovery Profiles

The above results focused on different segments of the response-recovery profiles, here the profiles are examined in their entirety (see Figures 1 and 4). When optimizing cluster number based on minimizing the ICL criterion, the clustering procedure identified 4 clusters for milk yield, BW, milk protein content, milk fat content in early lactation, BHB in early lactation, and insulin in early lactation. It identified 2 clusters for NEFA in early lactation and urea in late lactation; for the remaining traits, it identified 3 clusters. As can be seen from Figure 4, the clustering procedure generally performed well, classifying both differences that can be described as scaling differences (Figure 4a) and differences in shape (Figure 4b).

However, 2 issues should be noted; the first is that the cluster shapes are limited by the assumptions of the underlying piecewise model (i.e., a plateau value before the challenge and also 6 d postchallenge, a linear slope during the challenge, and a quadratic function for the recovery postchallenge). The second issue is that outlier animals often get assigned a class of their own (e.g., Figure 4a and 4c). This had the consequence that when the clustering was done using a fixed number of clusters (to allow comparison of cluster designation between traits the number was fixed to 3), the remaining variation between animals other than the outlier could be forced into only 2 clusters, resulting in a loss of type description. As shown in Figure 5, this was the case for milk protein content in early lactation, and to a lesser extent BHB and insulin also in early lactation. The issue of outlier animals should be further evaluated in studies with larger numbers of animals.

Because the clusters were numbered in ascending order of prechallenge level, it was possible to examine the extent of reranking of cluster designations between late lactation and the following early lactation period (whether the relative order of animals in cluster designations 1, 2, and 3 is changed). This was done in 2 ways, by visual inspection using the clustering without constraint on the cluster numbers, and using the ARI comparison with fixed cluster numbers. For milk yield, DMI, BW, and urea, relatively little reranking was observed and the corresponding ARI values were 0.3, 0.02, 0.41, and 0.33, respectively. (The range of ARI values across all comparisons for all traits was from 0 to 0.41). The apparent discrepancy between the ARI value for the late versus early lactation comparison for DMI relative to those of milk yield, BW, and urea was caused by the fact that although little reranking was found for DMI, 8 out of the 16 goats did decrease their class number by 1 (with a further 6 not changing class number). For milk yield, BW and urea, the numbers of

goats not changing class number were 10, 10, and 13, respectively. In contrast, for milk contents of fat and protein, and also for BHB, no clear association was found between late and early lactation class numbers (ARI values of 0.02, 0.05, and 0.04, respectively). For NEFA and glucose, these comparisons were not relevant because either the vast majority of goats were in 1 cluster (NEFA) or because an outlier goat skewed the cluster designation (glucose in late lactation). For insulin, 9 out of 16 goats kept the same rank (ARI = 0.1).

When comparing across traits, the highest ARI values in late lactation were for associations between milk yield and milk protein content (0.34), DMI and BW (0.23), DMI and glucose (0.21), milk contents of fat and protein (0.22), milk protein and BHB (0.39), and glucose and urea (0.33). The highest ARI values in early lactation were for associations between milk yield and DMI (0.19), DMI and BW (0.29), BW and insulin (0.28), and milk fat and BHB (0.19).

To provide a better visualization of the multivariate aspect, we used PCA on the cluster classes for all the

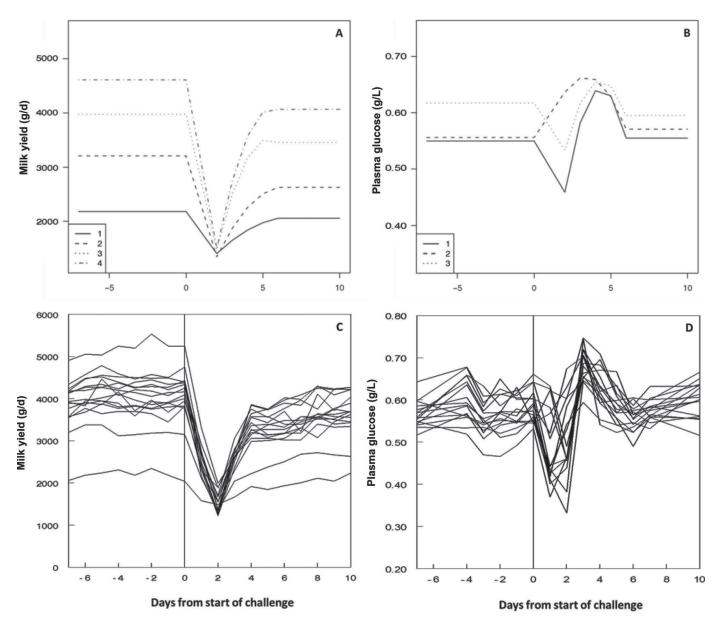


Figure 4. Classes of response-recovery profile for (a) milk yield and (b) plasma glucose in early lactation, generated by the combined piecewise mixed model and clustering procedure (see Materials and Methods). The 16 individual time trends from which these profile classes were derived are shown for comparison: (c) milk yield and (d) plasma glucose.

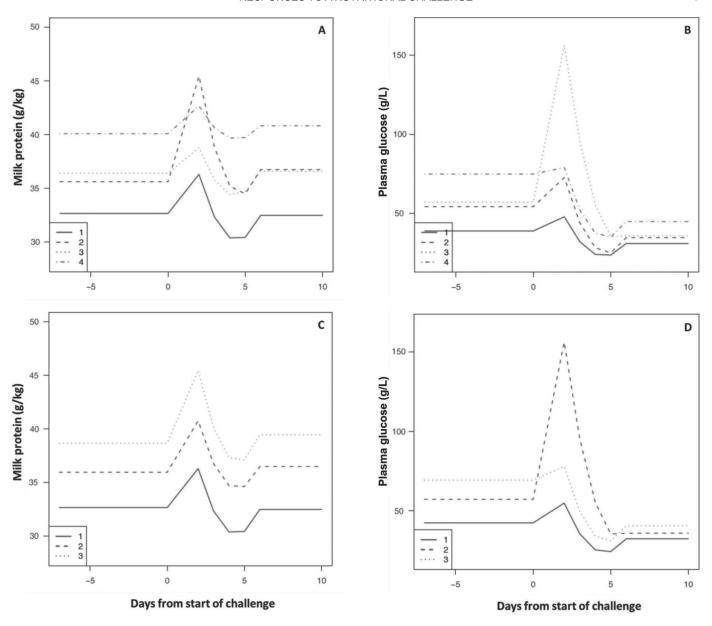


Figure 5. Comparison of profile classes based on variable cluster number assignment (a and b) with profile classes for the same measures based on a fixed cluster number (c and d). Two measures are shown: milk protein content (a and c) and plasma BHB (b and d) in early lactation. For milk protein content using the variable cluster number assignment (a), classes 1, 2, 3, and 4 contained 2, 8, 4, and 2 individuals, respectively. The 2 goats in class 1 combine with those in class 2 to give a new class 1 in the fixed cluster number assignment (c). For BHB, the variable cluster classes 2 and 4 (containing 5 and 3 goats) are merged into class 3 for 5 out of the 8 goats under the fixed cluster classes.

traits (Figure 6). In this PCA, within each trait, each cluster class was included as a separate variable [i.e., avoiding any a priori assumptions of relativity between cluster classes (linear or otherwise)]. For example, in Figure 6b it can be seen that DMI cluster classes 1, 2, and 3 fall in 3 distinct quadrants of the plot, which would not have occurred if DMI had been included as one variable with 3 levels (1, 2, and 3). The results of the PCA presented in Figure 6 are for a fixed cluster number across traits but were essentially the same

when using variable cluster numbers. From the loading plots for late and early lactation (Figures 6a and 6b), some patterns can be discerned. In both late and early lactation, the different cluster classes for milk fat content, milk protein content, and plasma NEFA are each positioned approximately on a linear transect, all running from the lower right to the upper left quadrant. In late lactation, the most closely correlated of these are NEFA and milk fat content, whereas in early lactation milk fat and protein contents are more closely

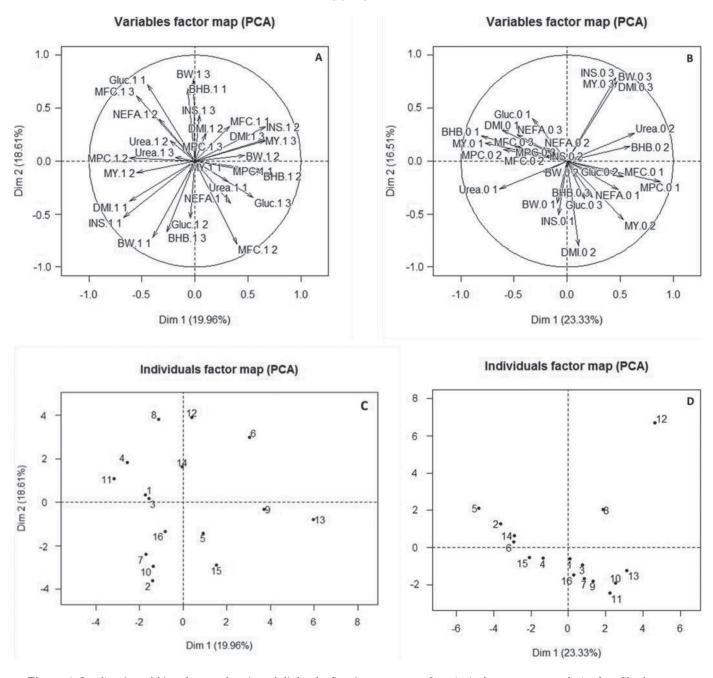


Figure 6. Loading (a and b) and score plots (c and d) for the first 2 components of a principal components analysis of profile classes across all measures in late lactation (a and c) and in early lactation (b and d). Within each measure, profile classes are considered as being independent from each other. Profile classes were derived from the combined piecewise fitting and clustering procedure (described in the Materials and Methods) using a fixed number of classes (3) across measures. To aid interpretation, the clusters were numbered, within lactation stage, in ascending order of prechallenge level. In the score plots, each goat is identified by a number, which is the same in both score plots.

associated. In late lactation, another transect can be discerned for DMI and milk yield, running from lower left to upper right, but in general no strong grouping is present of traits in the loadings plot, or of individual goats in the scores plot (Figures 6a and 6c, respectively). However, in early lactation, clearer positioning is present of traits into groups in the loadings plot

(Figure 6b). One of these groups (top right, Figure 6b) is essentially associated with only one animal (goat 12, Figure 6d). Within the 2 other groups, contrasts are present between milk yield and milk composition cluster numbers, and these groups lie roughly on the transect lines for plasma NEFA, milk fat, and milk protein.

DISCUSSION

General Response to the Nutritional Challenge

The overall effects of a 2-d nutritional challenge, replacing the standard TMR with straw, reported in this study using goats were broadly similar to other reports in the literature, although these come mainly from dairy cows. The decrease in milk yield in the present study (75% in late lactation and 62% in early lactation) was similar to the 57% decrease reported for black Moroccan goats that were feed deprived for 48 h (Hossaini-Hilali et al., 1993) and was equivalent to that found with a 2-d fast in dairy cows (McGuire et al., 1995; Agenäs et al., 2003). Using a more extreme feed restriction, 6 d of total fast, Reid et al. (1977) found that cows needed 49 d of recovery before their milk yield was normal. Toerien and Cant (2007) found that after 24 h severe restriction cows only needed 18 h of re-feeding before they were back on pretrial levels. In the present study, following 2 d on straw, goats regained prechallenge levels after approximately 3 d. In the recent study of Bjerre-Harpøth et al. (2012), where cows had a standard TMR diluted with 60% straw for 4 d, milk yield decreased by 40% and was largely recovered in 3 d. The study of Bjerre-Harpøth et al. (2012) also reported blood and milk metabolite changes for cows grouped according to early, mid, and late lactation. Allowing for the fact that the present challenge in goats was more severe, these 2 studies have general agreement with respect to magnitudes of changes in blood and milk traits, with a couple of exceptions. In the present study, milk protein content increased during the challenge, as did milk fat content, whereas in the study of Bjerre-Harpøth et al. (2012) only milk fat content increased with milk protein either decreasing (in early lactation) or not changing (in late lactation). Hossaini-Hilali et al. (1993) also found no increase in milk protein during feed deprivation. Further, in the present study and in the study of Hossaini-Hilali et al. (1993) urea increased markedly ($\times 2$) during the challenge, whereas it decreased in the study of Bjerre-Harpøth et al. (2012). These differences may be due to the degree of alteration in milk yield relative to the drop in protein supply, and the consequences of this on dilution within milk of the milk constituents. In the study of Bjerre-Harpøth et al. (2012), the CP content of the straw-diluted diet was adjusted to keep it above 12\%, which was not done in the present study. Otherwise, the prechallenge levels and responses in glucose and insulin were lower in early lactation than in late lactation in both studies, and the prechallenge levels and responses in NEFA and BHB were greater in early lactation than in late lactation in both studies. In

evaluating these results and the associated individual variation, it is important to remember that these are responses to a short-term challenge, which was useful for testing the methodology proposed in this study but may differ from responses to longer-term challenges.

Exploring Individual Variation in Response-Recovery Profiles

However, the main purpose of the present study was not to quantify the average effects of a nutritional challenge but rather to explore the individual variation in responses and recoveries, and to see if this variation could be used to identify and describe animal types. To do this we primarily used a simultaneous piecewise fitting and clustering procedure. With respect to characterizing individual differences in the profiles of measures prechallenge, during the challenge, and postchallenge, the piecewise mixed model performed well, given the number of parameters to estimate (4) relative to the number of measures in a profile (18). The same was true of the clustering procedure. Statistical analysis of the mixed model parameters revealed some important aspects for predicting responses and for repeatability. With respect to predicting responses, for both DMI and milk yield a significant linear effect was found of the prechallenge level on the slope, and thus the size, of the response (Table 2, Figure 3). Further, the responses in DMI and milk were very similar between early and late lactation when expressed as proportions of the prechallenge levels. This finding is in agreement with the results of Bjerre-Harpøth et al. (2012) for cows at different stages of lactation (49–273 DIM). No significant effects were found of prechallenge level on size of response for any of the other traits.

With respect to repeatability, significant coefficients were observed for the regressions between size of response in late lactation and in the following early lactation for DMI, milk protein content, BW, plasma BHB, and urea (and tendencies for milk yield and milk fat content; Table 3, Figure 3). Thus, for these traits, high responders in late lactation were also high responders in early lactation. This finding is highly promising for the characterization of adaptive capacity as it suggests that, despite the substantial homeorhetic changes associated with the transition from one lactation to another, these measures capture some of the innate, genetically driven, variability between individuals. This was not so evident for rate of deceleration in recovery (V_4) , where only plasma urea had a significant repeatability (data not shown). This quadratic component of the recovery was examined rather than the absolute recovery (i.e., combined linear and quadratic components) because

the differences in absolute recoveries would very largely reflect the amplitude of the response (correlation between linear slopes of response and recovery ranged from 0.49 to 0.99 for the different traits). However, examining partial components of recoveries is a relatively crude way of dealing with the correlations among the different phases of the time series (prechallenge, during the challenge, and postchallenge). Another, more elegant way to deal with these autocorrelations is to treat the time series in their entirety and examine their shapes for evidence of clusters that group different individuals. This was done using the methods described earlier permitting visualization of the different shape types, trait by trait (e.g., Figures 4 and 5). However, the use of the ARI statistic to examine correlations between cluster types was found not to be particularly informative. In our opinion, this is in part due to the loss of detail when standardizing the number of clusters per trait (to 3), a necessary condition for the ARI comparison. There is an argument for repeating this type of challenge study on a much larger number of animals to increase the statistical power of the cluster comparisons. Another shortcoming of the ARI method is that although it allows pairwise comparison across traits or lactation stages (in this study), it is still difficult to visualize across multiple traits.

Principal components analysis was used for a better visualization of the multivariate aspect. In late lactation a close correlation was observed between NEFA and milk fat content, whereas in early lactation milk fat and protein contents are more closely associated. A strong relationship between NEFA and milk fat content, or milk fat:protein ratio has previously been reported, especially in early lactation (Grieve et al., 1986; Moyes et al., 2013). The clearer positioning of traits in early lactation (Figure 6b) roughly on the transect lines for plasma NEFA, milk fat, and milk protein, and the fact that that the majority of individual goats are positioned along this transect (Figure 6d), is an important multivariate feature. It indicates a possible multivariate characteristic to describe differences between individuals using measures that relate to energy partition and metabolism. Interestingly, neither BHB nor glucose cluster classes exert a strong influence in principal components 1 or 2 in early lactation, although they are more evident in principal component 3 (not shown). The ability to describe differences between individuals in this way, and track these differences across stages of lactation, suggests that this may be a means to examine multivariate homeorhetic relationships.

The use of a simultaneous piecewise fitting and clustering procedure to provide the input variables (within trait cluster classes) to the PCA provides several

benefits: the variability that is captured describes the whole shape of the response/recovery, the clustering is independently done for each trait with no prior biological assumptions, and no relativity between cluster classes is assumed. The lack of prior grouping hypotheses implicit in this clustering method can be seen as an advantage of this approach relative to some of the indexing methods discussed below. Thus, the results of the PCA provide a simple means to begin to interpret the individual variability in the time-series profiles of response/recovery across multiple traits. Although the PCA visualization is useful for interpretation, this is clearly only one step if we wish to proceed toward methods to index differences in adaptive capacity. A follow-up step could be to derive an index directly from the PCA loadings. An example of this approach has been published for improving quantification of uterine health status from clinical measures (Gorzecka et al., 2011). Other approaches are to develop indexes directly from shape profiles without clustering [e.g., using Mahalanobis type statistics (Cohen et al., 2013)], or to use prior physiological knowledge to identify and combine candidate measures to, for example, index physiological imbalance (Moyes et al., 2013). However, not all of these examples may be suited to complex characteristics such as adaptive capacity that are by their nature emergent properties of multiple underlying mechanisms (Kitano, 2004). In this context, it is necessary to explicitly recognize the complex nature of the characteristic that we seek to index, and that we cannot directly measure it. What we have measured is the degree of perturbation resulting from a nutritional challenge, which is due to both the adaptive capacity of the animal and to the size of the challenge. The finding that the size of the response is related to the initial level of performance clearly indicates that the true size of challenge, to the internal milieu of the animal, varies from animal to animal even for an apparently constant externally applied challenge. Thus, it seems likely that the next step toward developing indexes for characteristics such as adaptive capacity is to build on multivariate results such as those presented here using models that explicitly invoke latent variables such as degree of perturbation and adaptive capacity. Such models are being developed in the time-series context and in simple form have even been applied to biological problems such as degree of infection (Højsgaard and Friggens, 2010).

CONCLUSIONS

Responses to nutritional challenge (2 d with access to straw only) can be characterized using piecewise mixed models and shape clustering. Dry matter intake and milk yield responses were found to be predictable from prechallenge levels, and a significant relation was found between responses in late lactation and in the following early lactation for DMI, milk protein content, BW, plasma BHB, and urea. Further, significant variation was observed between individuals in their response/recovery profiles, which suggests that defining adaptive capacity indexes from response/recovery profiles in multiple traits is possible.

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APPENDIX

Description of the Mixture Model Used for the Analysis of Response Recovery Profiles

Let $Y_j = (Y_{jl}, \ldots, Y_{jT})$ be the response profile of goat j. The (unknown) membership of goat j to one of the K classes is characterized through a label Z_j that takes values in $\{1, \ldots, K\}$. The response profile of a goat j that belongs to cluster k (i.e., for which $Z_j = k$), is

$$\begin{split} Y_{jt} &= V_1^k \times I_{(\mathbf{t} \leq 0)} + V_2^k \times t \times I_{(0 < \mathbf{t} \leq 2)} \\ &+ (V_3^k \times t + V_4^k \times t \times 2) I_{(2 < \mathbf{t} \leq 4)} + V_5^k \times I_{(t > 4)} + E_{jt} \end{split}$$

with the same notations as in the article. Parameters V_1^k, \ldots, V_5^k are similar to all goats belonging to class k but different between clusters. Such a model is called a mixture of linear models. The marginal distribution of profile Y_j is

$$\sum_{k=1}^{K} \pi_k \phi_k \left(Y_j \right), \tag{A1}$$

where π_k is the proportion of goats that belong to cluster k, and where ϕ stands for the (Gaussian) likelihood of profile Y_i according to class k.

The mixture model can be used to allocate each goat to a cluster. To do this, the probability of a goat to belong to cluster k given its profile is computed. This probability is called the posterior probability and is defined from equation [A1] by

$$\tau_{jk} = \frac{\pi_k \phi_k \left(Y_j \right)}{\sum_{l=1}^K \pi_l \phi_l \left(Y_j \right)}, \quad [A2]$$

The mixture parameters (proportions π_k , parameters V_1^k, \ldots, V_5^k , and the residual variance) are estimated using the EM algorithm. The EM algorithm is dedicated to the class of incomplete data models where the status of the observations is unknown. In this context, the missing data are the information on the cluster for each goat. The algorithm iterates 2 steps:

- In the E step, the parameters are fixed and the posterior probability for each observation to belong to each class is calculated using equation [A2].
- In the M step, the posterior probabilities are fixed the parameters of each class are estimated using a weighted regression, in which the weights are given by the posterior probabilities.

Once convergence is achieved, one can classify an animal into a cluster using the maximum a posteriori rule: one just classifies the animal according to its highest posterior probability.

The whole strategy requires the number of classes to be known, something unrealistic in practice. When K is unknown, the model can be fit for different potential values of K, and the best value is then selected using a penalized criterion. In this article, we considered the classical integrated completed likelihood (Biernacki et al., 2000). This criterion is a penalized likelihood criterion, which tends to select a parsimonious number of relatively homogeneous clusters.